

STAT 5203 Lecture 3

Admin

1. Monday office hours (Before 6)
2. HW2 due date (22nd or 24th)
3. Videos on birthday problem and matching problem

Important dates

Sept 15 : HW1 due

Sept 24 : HW2 due

Sept 29 : Midterm

Agenda

1. Recap on conditional probability + law of total probability
2. Bayes rule
3. Independent events
4. Random variables
5. Intro to pdfs and cdfs

Recap: Conditional prob $A, B \subseteq S$

$$\begin{aligned} P(A) > 0 \quad P(B|A) &= \frac{\overset{\text{Conditional prob}}{P(AB)}}{P(A)} \Rightarrow P(AB) = P(B|A)P(A) \quad \begin{matrix} \text{Chain rule} \\ \text{mult. rule} \end{matrix} \\ &\Rightarrow P(B) = P(B \cap A) + P(B \cap A^c) \\ &= P(B|A)P(A) + P(B|A^c)P(A^c) \\ &\quad \swarrow \text{Law of total prob.} \end{aligned}$$

Today: Bayes rule "flip conditional probabilities"

Setup: Suppose we want $P(B|A)$, but we only know $P(A|B)$, $P(B)$, ...

$$P(B|A) = \frac{P(AB)}{P(A)} = \frac{P(A|B)P(B)}{P(A|B)P(B) + P(A|B^c)P(B^c)} \quad \begin{matrix} \text{Bayes} \\ \text{rule} \end{matrix}$$

Example: A lab test is 95% correct in detecting a disease when it is present. The false positive rate is 1%. If 0.5% of the population has the disease, what is the prob that someone has the disease given that they test positive?

D = the person has the disease

P = the person tests positive

$$0.95 = P(P|D)$$

$$0.01 = P(P|D^c)$$

$$0.005 = P(D) \quad P(D^c) = 0.995$$

$$P(D|P) = \frac{P(P|D)P(D)}{P(P|D)P(D) + P(P|D^c)P(D^c)} = \frac{0.95 \times 0.005}{0.95 \times 0.005 + 0.01 \times (1 - 0.005)}$$

"with probability"

Example: A student is taking a multiple choice exam. Student knows the answer w.p. p and they guess w.p. $1-p$. If the student guesses, they get the question correct w.p. $1/m$ ($m = \#$ of choices).

Suppose a student gets the answer correct. What is the prob that they guessed?

A = student knows answer

B = student gets answer correct

General Bayes formula: If $F_1, \dots, F_n$ are disjoint and $\bigcup_{i=1}^n F_i = S$, then for some $E \subseteq S$,

$$P(F_i | E) = \frac{P(E | F_i) P(F_i)}{\sum_{j=1}^n P(E | F_j) P(F_j)}$$

$P(F_i)$ "prior probability"
 $P(F_i | E)$ "posterior probability"

Example: A plane is missing and has been assumed to have crashed in one of 3 regions (R_1, R_2, R_3) each with equal probability. Assume that if the plane crashed in R_1 , the prob. of finding it is 0.8.

What is Prob that the plane crashed in R_i ($i=1, 2, 3$) if the search of R_1 is unsuccessful?

R_i = plane is in region i $P(R_i) = 1/3$ $i=1, 2, 3$

E = region 1 search was unsuccessful

$$P(R_1 | E) = \frac{P(E | R_1) P(R_1)}{\sum_{j=1}^3 P(E | R_j) P(R_j)} = \frac{0.2 \times \frac{1}{3}}{0.2 \times \frac{1}{3} + 1 \times \frac{1}{3} + 1 \times \frac{1}{3}} = \frac{1}{11}$$

$$P(R_2 | E) = P(R_3 | E) = \frac{5}{11} \quad (\text{exercise})$$

	$P(R_i)$	$P(R_i E)$
$i=1$	$1/3$	$1/11$
$i=2$	$1/3$	$5/11$
$i=3$	$1/3$	$5/11$

Independent Events

DEF: $A, B \subseteq S$, A and B are independent if $P(A \cap B) = P(A)P(B)$

* Equivalently, if $P(A) > 0$, $P(B) > 0$. Then A and B are indep if

$$P(A|B) = P(A)$$

$$P(B|A) = P(B)$$

* (HW) If $A \cap B = \emptyset$, then they are not independent

Other facts: suppose A, B indep.

Then, ① A, B^c indep. want to prove that $P(AB^c) = P(A)P(B^c)$

② A^c, B indep.

③ A^c, B^c indep.

Independence of multiple events: $A_1, A_2, \dots, A_n$

DEF: $A_1, A_2, \dots, A_n$ are pairwise indep. if $\forall i, j \in \{1, \dots, n\} \ i \neq j$
 $P(A_i A_j) = P(A_i)P(A_j)$

DEF: $A_1, A_2, \dots, A_n$ are mutually indep if $\forall$ subsets of events
 $A_{i_1}, A_{i_2}, \dots, A_{i_r} \ 2 \leq r \leq n$, we have

$$P(A_{i_1} A_{i_2} \dots A_{i_r}) = P(A_{i_1}) \times \dots \times P(A_{i_r})$$

e.g., A_1, A_2, A_3, A_4, A_5 . Then

$$P(A_1 A_2 A_3) = P(A_1)P(A_2)P(A_3)$$

$$P(A_1 A_2) = P(A_1)P(A_2)$$

$$P(A_1 A_2 A_3 A_4) = P(A_1)P(A_2)P(A_3)P(A_4)$$

$\vdots$

Mutual independence $\Rightarrow$ pairwise indep.

$\nLeftarrow$

e.g., tossing 2 coins
rolling 2 dice

$$S = \{HH, HT, TH, TT\}$$

$$P(HH) = \frac{1}{4} = \frac{\# HH}{\text{total}}$$

$$P(HH) = P(H) \times P(H) = \frac{1}{2} \times \frac{1}{2} = \frac{1}{4}$$

Repeated independent trials

e.g., toss two coins $P(H) = \frac{1}{2}$
2 trials

e.g. Roll a dice 5 times. What is the prob. of rolling a 5 exactly 2 times?

$$E_i = \text{roll a 5 on the } i^{\text{th}} \text{ trial} \quad P(E_i) = \frac{1}{6} \quad P(E_i^c) = \frac{5}{6}$$

$$P(E_1 E_2^c E_3^c E_4^c E_5^c) = P(E_1)P(E_2^c)P(E_3^c)P(E_4^c)P(E_5^c) = \left(\frac{1}{6}\right)^2 \left(\frac{5}{6}\right)^3$$

$S \ S \neq S \neq S \neq S$

$S \neq S \neq S \ S \neq S$

$\vdots$

$$P(2 \text{ 5's}) = \underbrace{\# \text{ of ways to get 2 5's in 5 rolls}}_{\text{independent } \binom{5}{2}} \times \underbrace{\left(\frac{1}{6}\right)^2 \left(\frac{5}{6}\right)^3}_{\text{success in trial } i}$$

In general, if we have N trials and in each trial $P(E_i) = p$
 $P(K \text{ success}) = \binom{N}{k} p^k (1-p)^{N-k} \leftarrow \text{Binomial distribution}$

Useful : If $A_1, \dots, A_n$ are indep. and $P(A_i) = p_i$

$$\begin{aligned} P(\text{at least one } A_i \text{ occurs}) &= 1 - P(\text{none of } A_1, \dots, A_n \text{ occur}) \\ &= 1 - (1-p_1) \times (1-p_2) \times \dots \times (1-p_n) \end{aligned} \quad \left. \vphantom{\begin{aligned} P(\text{at least one } A_i \text{ occurs}) &= 1 - P(\text{none of } A_1, \dots, A_n \text{ occur}) \\ &= 1 - (1-p_1) \times (1-p_2) \times \dots \times (1-p_n) \end{aligned}} \right\} \text{ Birthday Problem}$$

e.g., $P(m \text{ failures before } n \text{ success}) = \binom{m+n-1}{m} p^n (1-p)^m$ Negative Binomial

Ch3 DeGroot : Random variables, distributions, pmf, pdf, cdfs
Random variables outcomes $\rightarrow$ functions of outcomes

e.g., toss 2 coins $S = \{HH, HT, TH, TT\}$
of heads for

HH	2
HT, TH	1
TT	0

$$P(1 \text{ head}) = \frac{1}{2}$$

DEF : A random variable, X , is a function from the sample space to the real numbers $X: S \rightarrow \mathbb{R}$

$X = \# \text{ of heads}$
 $X: S \rightarrow \{0, 1, 2\}$

$S = \{HH, HT, TH, TT\}$
↓ ↓ ↓ ↓
2 1 1 0
$\in \mathbb{R} \in \mathbb{R} \in \mathbb{R} \in \mathbb{R}$

$$P(X=0) = \frac{1}{4}$$

$$P(X=1) = \frac{1}{2}$$

$$P(X=2) = \frac{1}{4}$$

Random Variable

→ Connection between r.v. and probabilities

More generally, $E \subset \mathbb{R}$ Then $P(X \in E) = P(\{s \in S : X(s) \in E\})$
e.g., $P(X=1) = P(\{HT, TH\}) = \frac{1}{4} + \frac{1}{4} = \frac{1}{2}$

e.g., toss two coins. Start with \$1. If H then double our fortune
T then lose everything

$X = \text{fortune after 2 flips}$

$S = \{HH, HT, TH, TT\}$

↓ ↓ ↓ ↓
4 0 0 0

$$P(X=0) = \frac{3}{4}$$

$$P(X=4) = \frac{1}{4}$$

$$P(X=a) = 0 \text{ for } a \neq 0, 4$$

DEF: The distribution of X is the collection of all probs. of the form $P(X \in C) \forall C \subseteq \mathbb{R}$ such that $\{X \in C\}$ is an event

Discrete random variables: "There are a countable number of possible values of X "

DEF: (Probability mass function or pmf). For a discrete random variable, the pmf $p: \mathbb{R} \rightarrow [0, 1]$ of X is given by

$$p(x) = P(X = x) = P(\{s \in S : X(s) = x\})$$

↑ realization of X
Random variable

pmf properties: X can have values $x_1, x_2, x_3, \dots$. The pmf satisfies

- (1) $p(x_i) \geq 0 \quad \forall i = 1, 2, 3, \dots$
- (2) $p(x) = 0$ for $x \notin \{x_1, x_2, \dots\}$
- (3) $\sum_{i=1}^{\infty} p(x_i) = 1$

common pmfs

(1) Binomial distribution: n indep. each with success prob. p

$$P(X = k) = \binom{n}{k} p^k (1-p)^{n-k} \quad \text{for } k = 0, 1, \dots, n$$

$$P(\underbrace{\text{Number of success}}_X = k)$$

$$p(k) = \begin{cases} \binom{n}{k} p^k (1-p)^{n-k} & \text{for } k = 0, 1, 2, \dots, n \\ 0 & \text{otherwise} \end{cases}$$

Note: define pmf for all possible values

(1a) Bernoulli distribution: Binomial dist. with $n=1$

$$p(k) = \begin{cases} 1-p & k=0 \\ p & \text{for } k=1 \\ 0 & \text{o/w} \end{cases}$$

(2) Geometric distribution: indep. each with success prob. p

$X = \#$ of trials required for first success

$$P(X = k) = (1-p)^{k-1} p$$

$$p(k) = \begin{cases} (1-p)^{k-1} p & k = 1, 2, \dots \\ 0 & \text{o/w} \end{cases}$$

(3) Negative binomial

Example: The pmf of a r.v X is given by

$$p(i) = \begin{cases} c \frac{L^i}{i!} & \text{for } i = 0, 1, 2, \dots \\ 0 & \text{o/w} \end{cases}$$

c is a constant
 $L > 0$

(a) Find the value of c

$$\sum_{i=0}^{\infty} \frac{c L^i}{i!} = 1 \Rightarrow c \underbrace{\sum_{i=0}^{\infty} \frac{L^i}{i!}}_{e^L} = 1 \rightarrow c e^L = 1 \Rightarrow c = e^{-L}$$

(Solve for c)

(b) $P(X=0) = c = e^{-L}$ $\frac{e^{-L} L^0}{0!} = e^{-L}$

(c) $P(X > 2) = \sum_{i=3}^{\infty} \frac{e^{-L} L^i}{i!}$

(easier) $= 1 - P(X=1) - P(X=0)$
 $= 1 - \frac{e^{-L} L^1}{1!} - e^{-L}$

Remark: The above pmf is the pmf of the Poisson distr. w/ param L

DEF: (Cumulative distribution function or cdf) The cdf $F: \mathbb{R} \rightarrow [0, 1]$ of an r.v X is defined as follows $F(b) = P(X \leq b)$ for $b \in \mathbb{R}$

Note: If X is discrete, $F(b) = \sum_{x \leq b} p(x)$

e.g., flip 2 coins. $X = \#$ of heads

$P(x) = \begin{cases} 1/4 & x=0 \\ 1/2 & x=1 \\ 1/4 & x=2 \\ 0 & \text{o/w} \end{cases}$	$F(x) = \begin{cases} 0 & (-\infty, 0) \\ 1/4 & [0, 1) \\ 3/4 & [1, 2) \\ 1 & [2, \infty) \end{cases}$	$\begin{matrix} p(0) \\ p(0) + p(1) \\ p(0) + p(1) + p(2) \end{matrix}$
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$$F(1) = P(X \leq 1) = P(X=0) + P(X=1) = \frac{1}{4} + \frac{1}{2} = \frac{3}{4}$$

$$F(1.5) = P(X \leq 1.5) = P(X=0) + P(X=1) = \frac{3}{4}$$

$$F(3) = P(X \leq 3) = p(0) + p(1) + p(2) = 1$$